Cancer Mortality Among a Representative Sample of Nonsmokers in the United States During 1966-68

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ABSTRACT-Data are presented on cancer and total mortality among a representative sample of nonsmokers and the total population 35-84 years of age in the United States during 1966-68 that measured the influence of cigarette smoking on mortality rates, independent of other health-related factors. Of all U.S. white males, those who never smoked digarettes have a total ace-adjusted cancer death rate which is 37% less than that of males as a whole and 53% less than that of those who currently smoke cigarettes. Correspondingly, of all U.S. white females, those who never smoked cigarettes have a total age-adjusted cancer death rate which is 15% less than females as a whole and 33% less than that of those who currently smoke cigarettes. The largest cancer rate reduction in the nonsmokers is concentrated In the respiratory system. Nonsmokers have an age-adjusted total death rate which is about 20% less than the population as a whole and about 43% less than current cigarette smokers. These and other results and methodologic issues are discussed.—JNCI 65: 1175-1183, 1980.

Previous studies have established a strong, positive correlation between cigarette smoking and mortality, with cancer mortality showing some of the highest correlations. Beginning in 1964, the Smoking and Health Report of the Surgeon General has summarized the evidence (1). Only one previous study in the United States has been based on representative national samples and that was limited to lung cancer deaths during 1958-59 (2, 3).

The findings here are based on a 1966-68 survey that measures among a representative U.S. sample the effect of smoking status on all causes of death independent of other health-related factors. The nonsmokers in this survey are meant to be used as a reference for comparisons with other nonsmoking populations who may be subject to additional health-related influences and selection biases. Our focus is on cancer mortality among nonsmokers and their risks relative to those experienced by the total population. "Nonsmokers" here are defined as persons who smoked fewer than 5 packs of cigarettes (100 cigarettes) in their lifetimes. However, they include former and current cigar and/or pipe smokers.

Unlike most other low-risk populations, nonsmokers comprise a substantial proportion of the adult population. They are not limited to a particular geographic region, ethnic group, or religion. Moreover, our use of national probability samples to examine the mortality experience of nonsmokers minimizes selection biases with respect to other environmental, socioeconomic, and life-style factors.

MATERIALS AND METHODS

Investigators who conducted most previous studies of cigarette smoking and mortality used a single-sample, prospective design. Mortality ratios were obtained from the relative proportions of the original sample who died within a specified follow-up period and then were weighted by survival times. However, serious disadvantages associated with representativeness and problems of follow-up are evident in the prospective studies (1). Nonresponse may have been as much as \$2% overall and even greater among smokers. The bias in the mortality ratios of the respondents was estimated to be about a 0.2 or 0.3 overstatement because of lower mortality ratios of the nonrespondents. The low level of nonresponse in the present study (5%) indicated that nonresponse was negligible as a source of bias.

Our approach was to use a cross-sectional, two-sample design to estimate the relative risk of non-smokers' mortality. This study was analogous to the one of representative samples of 1958-59 lung cancer deaths among U.S. whites (2, 3). One sample was representative of U.S. deaths during 1966-68, and the other resembled the general U.S. population during the same period. The smoking characteristics of both samples were determined and then mortality rates were calculated as a function of smoking status. Some limited results of the study, described in detail elsewhere, have been published (4-6).

Sources of data.—Data were from two sources. Estimates of observed deaths of nonsmokers 35-84 years of age in the United States in 1966-68 were based on the NMS, a follow-back survey linked to a probability sample of 19,526 death registration records that in-

ABBREVIATIONS USED: CPS=Current Population Survey(s); ICD=International Classification of Diseases; NCHS=National Center for Health Statistics; NMS=National Mortality Survey; SMR=standardized mortality ratio(s).

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cluded deaths of 11,318 white males and 5,636 females.5 The underlying cause of each death was assigned in accordance with the Seventh Revision of the ICD (7, 8). Questionnaires were mailed to surviving family members and others named on the death certificate who provided smoking histories and social characteristics of the deceased. The second source of data is the CPS conducted by the U.S. Bureau of the Census in August 1967 (9). Smoking and socioeconomic information comparable to that for the decedent was obtained by means of household interviews for a probability sample of 60,920 adults 35-84 years of age, including 25,266 white males and 29,308 white females.4 The technical details of each survey are given in the Appendix. Also included there are the age-sex distributions of the deaths and population-at-risk by smoking status and a sample calculation of the weighted number of deaths when the sample number of deaths is used.

Inferences that can be drawn from the data in this study are conditioned by certain features of its design, sampling, and data-collection methods, as discussed in detail in the Appendix. In prospective studies, the decedents are a subset of the total study population; examples are the American Cancer Society and U.S. Veterans cohort studies in which a beginning sample is initially questioned regarding both current and past smoking habits (10, 11). The decedents of a prospective study, who are in the original sample, match exactly with an individual in the study population. The cross-sectional design assumes statistical matching between the two independently drawn samples: decedents and population. This problem is common to conventional vital statistics rates and must be balanced with the expense, time, and follow-up problems of prospective studies. The value of a two-sample crosssectional design was demonstrated in national studies of smoking and lung cancer mortality during 1958-59

Definitions.—Detailed estimates of cigarette smoking history by amount smoked and other relevant smoking variables were not made. Instead, persons were simply classified by lifetime history of having "never smoked" or "ever smoked" cigarettes (at least 5 packs or 100 cigarettes). "Current" cigarette smokers are persons who have smoked at least 100 cigarettes in their lifetimes and have smoked cigarettes during the past 12 months. We coordinated smoking questions in the supplement to the CPS with those in the death follow-up survey in an effort to gain comparability. Unlike sizable proportions of the CPS sample who provided data on their smoking habits and other personal characteristics, such data were not available for the NMS death sample. Moreover, a mailed questionnaire was used in the NMS, whereas a household interview was used by the CPS for collection of data on

the living population. Evidence from another survey indicated that agreement between the smoking his reported for a decedent by next of kin and by his in reports when alive was about 90%. Moreover, Haenszel et al. (2) noted that this type of bias is largest with respect to medium and heavy smokers but virtually nonexistent for nonsmokers.

RESULTS

Table 1 contains the average annual age-specific total death rates for 1966-68 for U.S. white males and . females by smoking status: never smoked cigarettes, currently smoke cigarettes, and the total sample. The rates are given in 10-year intervals for adults 35-84 years old, with an overall rate age adjusted to the 1940 . U.S. population. The 1940 U.S. population is the standard reference used by the NCHS (7). Included in table I are age-specific total death rates for all U.S. whites based on the NMS and CPS and the average total death rates as published in the 1966-68 and 1970 Vital Statistics of the United States (8). The NMS-CPS rates generally agree with the 1966-68 U.S. rates to within 3%, except in the 75- to 84-year-old category, in which they are up to 10% higher than the U.S. rates. (The CPS includes only the noninstitutionalized population as explained in the Appendix.) Agreement between the survey rates and the U.S. vital statistics is good. Finally, it is evident that the nonsmokers in this survey have significantly lower rates than those of t total population, which conforms with previous ex demiologic studies of nonsmokers (1, 10, 11). Nonsmokers have an age-adjusted total death rate about 20% less than that of the population as a whole and about 43% less than the rate for current cigarette smokers.

In addition to the 1966-68 average annual agespecific cancer death rates for U.S. white males and females by smoking status, table 2 includes the rates for total cancer and for cancers of the lung, breast, and prostate, all of which are based on more than 100 sample deaths. The U.S. white males who never smoked cigarettes have a total age-adjusted cancer death rate 37% less than that of U.S. white males as a whole and 53% less than that of white males who are currently eigarette smokers. With the rates for total cancer based on the NMS and CPS are those given in . the 1966-68 and 1970 U.S. vital statistics. The agreement between the survey and the vital statistics rates is within 11% for those less than 65 years old. Over 65 years, the survey rates are 7-17% higher than those for . vital statistics. These differences are greater than the corresponding differences for total mortality primarily because of statistical fluctuation in the sampling of cancer deaths in the NMS.

The 1966-68 weighted numbers of observed cancer deaths for U.S. whites who never smoked cigarettes for the sites listed based on at least 10 sample deaths in males and/or females are given in table 3. Also included are the expected numbers of cancer deaths

³ National Center for Health Statistics, 1966-68 National Mortality Survey. Unpublished technical notes and computer tapes, 1976.

	U.S. whites	U.S. whites		All U.S. whites	
Age, yr	who never smoked cigarettes*	who cur- rently smoke cigarettes	1966-68 NMS + CPS*	NMS+ vital	
		Ma	les		
85-44 45-54 55-64 65-74 75-84 Total: 35-84	216 570 1,686 3,654 9,528 1,425	442 1,183 2,878 6,572 13,930 2,459	352 937 2,352 4,933 10,749 1,918	842 904 2,246 4,968 9,918 1,852	844 883 2,203 4,810 10,099 1,827
		Fem	ales		
85-44 45-54 55-64 65-74 75-84 Total: 85-84 ^b	152 833 771 2,398 7,044 894	245 597 1.444 4,129 13,123 1,606	201 464 1,015 2,702 7,600 1,058	195 462 1.022 2.628 6.890 1,017	193 463 1,015 2,471 6,699 989

Death rates are based on the 1966-68 NMS and August 1967 CPS.

Rates are age adjusted by the direct method to the 1940 U.S. population 85-84 yr old.

based on the death rates among all the whites in the 1966-68 NMS sample. With this group as the reference population, the total cancer SMR for whites were 67 and 87% for male and female nonsmokers, re-

spectively. Using 1970 U.S. whites as the reference population, as was done in a previous publication summarizing some of these NMS results (5), we determined that the total cancer SMR were 70% for white

TABLE 2.—The 1966-68 average annual age-specific cancer death rates (deaths/100,000) among U.S. whites by age, sex, and cigarette emoking status"

		Lung cancer			Total cancer						
Age, yr	U.S.	U.S. whites		U.S whites who	U.S. whites	U.S. whites	All U.S. whites				
	who never	who currently smoke cigarettes	All U.S. whites	never smoked cigarettes	who never smoked cigarettes	who cur- rently smoke cigarettes	1 966-68 NMS + CPS'	1966-68 U.S. vital statistics	1970 U.S vital statistics		
					Males						
35-44	2.3 (2)	17.5 (38)	12.9 (48)	0.7 (1)	32.0 (28)	56.8 (125)	49.2 (185)	50	50		
45-54	3.5 (3)	81.1 (163)	59.3 (213)	2.1(3)	80.7 (42)	211.2 (340)	172.4 (476)	170	172		
55 - 64	33.4 (24)	270.4 (365)	177.6 (493)	15.2 (18)	306.0 (102)	685.7 (646)	527.5 (946)	479	498		
65-74	63.8 (43)	533.9 (298)	288.5 (499)	79.9 (89)	655.5 (204)	1,531.3 (543)	1,079.7 (1,041)	985	997		
75-84	87.9 (36)	746.1 (122)		244.4 (165)	1,360.8 (294)	2,224.3 (262)	1,548.0 (760)	1,472	1,593		
Total: 35-84'	19.9 (108)	179.8 (986)		24.0 (276)	232.8 (670)	497.4 (1,916)	368.6 (3,408)	84 3.9	854.9 		
				1	Females						
35-44	0.4 (1)	8.6 (16)	4.2 (19)	23.8 (46)	56.7 (110)	80.1 (134)	72.9 (288)	6 5	62		
45-54	2.8 (6)	25.4 (43)	15.3 (65)	45.6 (46)	133.2 (141)	197.2 (179)	172.1 (377)	175	177		
55-64	10.9 (24)	51.9 (46)	25.4 (85)	85.9 (45)	268.5 (166)	412.6 (141)	335.7 (360)	3 30	3 39		
65-74	18.7 (37)	103.7 (31)	34.3 (82)	121.2 (28)	576. 9 (182)	877.1 (66)	627.1 (274)	5 65	55 5		
75-84	48.4 (55)	143.9 (9)	55.5 (68)	143.0 (19)	921.8 (190)	1,303.9 (21)	999.2 (229)	868	904		
Total: 35-84'	7.5 (123)	89.4 (145)	17.5 (319)	59.4 (184)	220.9 (789)	328.8 (541)	260.6 (1,528)	244.5	246.2		

Unweighted No. of sample deaths are given in parentheses.

Death rates are based on 1966-68 NMS and August 1967 CPS. In column 5, values are for prostate cancer in males and breast cancer in females.

'Rates are age adjusted by the direct method to the 1940 U.S. population 35-84 yr old.

TABLE 8.—The 1966-68 weighted No. of observed (Obs) and expected (Exp) deaths and SMR by major cancer site for a representative sample of U.S. whites \$5-84 years old who never smoked eigarettes

Primary site	ICD (Seventh	9,8	25,416 ma	le nonsmoke	278	28,9	45,780 fem	ale nonsmo	kers
	Revision) code No.	Obs	Sample No.	Exp	5MR'	Obs	Sample No."	Exp	SMR'
Total malignant neoplasms	140-205	106,478	670	157,798	67	237,261	789	071 100	
Buccal cavity and pharynx	140-148	2,728	56	4,403	62	-		271,588	87
Digestive organs and peritoneum	150-159	45,029	135	48,457	93	1,702	28	8,248	52
Esophagus	150	1,782	87	8,216	55	74,968	180	87,580	86
Stomach	151	7,920	17	10.651		1,601	- 26	1,992	80
Colon	153	17.967	44		74	10,553	27	13,325	79
Rectum	154	7,560	18	13,553	133	80,438	61	82,446	94
Liver, galibladder, and biliary	155-156	3,282	6	6,463	117	8,117	18	10,573	77
passages		U1404	v	5,272	62	11,717	22	13,311	88
Pancreas	157	4,759	9	7 070	••				
Respiratory system	160-165	18,038	115	7,870	6 0	11,658	23	13,983	83
Lung, bronchus, and trachea	162-163	8,733	108	46,581	28	12,908	131	21,711	59
Breast	170	0,100	100	89,991	22	9,079	123	17,122	5 3
Female genital system	171-176					55 ,051	184	58,659	94
Cervix	171					38,385	124	41,576	92
Corpus uteri	172					7,790	80	11,642	67
Uterus, NOS"	173-174					1,126	6	1,937	58
Ovary, fallopian tube, and	175					5,348	16	6.316	85
broad ligament	110					21,563	68	19,393	111
Male genital system	177-179	13.784	070					, -	
Prostate gland	177		278	16,563	83				
Urinary system	180-181	13,534	276	16,263	83				
Nervous system	193	5,881	13	9,578	61	6,816	12	7.926	86
Lymphomas	200-202	2,478	10	2,609	95	6,830	24	5.543	123
Leukemia	204	4,187	14	6,417	65	7,991	21	9,131	88
		3,361	12	5,238	64	8,754	24	10,107	87
Total malignant neo- plasms'	140-205	106,478	670	152,037"	70	237,261	789	250,521	95

[&]quot; Sample No. of deaths from which weighted No. of observed deaths is calculated with the use of poststratified ratio estimation procedure (see text).

NOS = not otherwise specified.

male and 95% for female nonsmokers. For male nonsmokers, the lowest SMR values occurred for the traditional smoking-related cancer sites of the respiratory system, buccal cavity and pharynx, esophagus, pancreas, and urinary system, with a combined SMR of 39%, whereas the remaining cancer sites had a combined SMR of 91%. However, the lowest SMR values for female nonsmokers occurred for the cancer sites of the lung and buccal cavity and pharynx, with a combined SMR of 53%; the SMR values for most of the other sites were between 80 and 95%. Because of the complex sampling and weighting procedures used in these data, confidence limits on an individual SMR cannot be precisely calculated, although approximate limits can be determined with the use of the sample number of deaths, if one assumes they are subject to a Poisson distribution (2).

DISCUSSION

The 1966-68 NMS represents the first and only attempt to date to determine death rates for all causes of death among a representative sample of the U.S. popula-

tion classified according to smoking status. The basic findings in this paper show that during 1966-68, U.S. whites 35-84 years old who never smoked cigarettes have age-adjusted death rates from cancer and all causes substantially below those of the population as a whole and only about one-half the corresponding rates of those who currently smoke cigarettes. Phrased in other terms, whites who never smoked cigarettes have overall SMR which are substantially less than 100%. Specific comparisons vary by a few percentage points with different reference populations or with direct versus indirect standardization, but the basic patterns remain the same. Furthermore, these results indicate that most of the reduction in cancer deaths among male nonsmokers occurs among the traditional smoking-related cancer sites of the respiratory system, buccal cavity and pharynx, esophagus, pancreas, and urinary system. The single greatest reduction in cancer deaths among female nonsmokers occurs in the respiratory system; however, small reductions also occur at many other sites.

These results provide a methodologically sound data set on which to base comparisons with other popula.

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Expected No. of deaths from cancer is based on 5-yr, age-specific death rates determined for U.S. whites (1966-68) from the ta 1966-68 NMS and August 1967 CPS (see text).

SMR is the percentage defined as Obs+Exp × 100.

Expected No. of deaths from cancer is based on 5-yr, age-specific death rates for U.S. whites in 1970 (8).

tions at low risk to cancer. In other words, any nonsmoking population who has death rates similar the 1966-68 NMS nonsmokers is probably at low risk solely because they do not smoke. These results help delineate the specific influence of cigarette smoking on mortality rates from other health-related factors.

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APPENDIX

National Mortality Survey

The NMS is based on a systematic random sample of all deaths occurring and registered in the United States.3 Death registration is essentially 100% complete and is performed on a state-of-occurrence basis with the use of the official State death certificate. The States generally follow the federally recommended model death certificate. Date, place, and cause of death, age, sex, and other demographic characteristics are important legal facts noted on the document. One can extend the research potentials of death certificates by using them as a sampling frame for surveys (such as the NMS) in which additional information is collected through informants, funeral directors, physicians, and others named on the certificate.

The scope of the NMS is all deaths of persons 35-84 years old that occurred in the United States during 1966, 1967, or 1968. The sample was stratified for month of death, geographic area, age, and selected causes of death related to smoking. The total sample of 19,526 decedents between 35 and 84 years of age was drawn in two stages. 1) A 10% sample is represented in the monthly "Current Mortality Survey" of death certificates submitted to the NCHS on microfilm by all States, the District of Columbia, and other death registration areas. 2) The strata of the 10% sample are subsampled; each 26th case was taken on the average. The overall sampling rate is therefore 1/260. The demographic and medical information coded from the sampled death certificate is supplemented by information from a mailed questionnaire on smoking habits, marital status, income, and other characteristics of the deceased. One-half of the respondents are spouses or ex-spouses of the deceased, one-fourth are either parents or offspring, one-eighth are siblings, and the remaining 10% are comprised of more distant relatives, friends, neighbors, and other associates. The overall response rate of 95% in the NMS is comparable to that of the

After follow-up queries of questionnaires with incomplete, inconsistent, or otherwise inadequate responses, smoking information was still missing for about 8% of the sample deaths. For these cases, smoking history was assigned by computer processing of the records of the decedents. This procedure, in the absence of better information, assumes that the characteristics for which reported information is missing are like those of demographically similar individuals for whom the information in question was reported.

This processing was performed by the NCHS, which also supplied the magnetic tape of 1966-68 NMS records. Each record combines the smoking and other NMS questionnaire information with that from the death certificate. In addition, each record contains an inflation factor (sample weight) that reflects the sampling fraction applicable to the decendent's demographic cause-of-death stratum. This sample weight is determined by a poststratified ratio estimation procedure which forces the total number of deaths estimated from sampling to agree with the total number of deaths from which the sampling was done. The weight is a function of age, year of death, and cause of death and varies from about 10 to 640. Appendix table 1 shows an example of how the total weighted number of 55,051 deaths is calculated from the 184 sample deaths due to breast cancer among U.S. white females who never smoked cigarettes. The age-sex distributions of the total weighted deaths from all causes according to smoking status are given in Appendix table 2. The underlying cause of death (item #18 of the death

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APPENDIX TABLE 1.—Sample calculation of 1968-68 average annual age-specific and age-adjusted death rates for breast cancer for U.S. white females who never smoked cigarettes by the poststratified ratio estimation procedure

Age, yr	Unweighted No. of sample deaths by yr				Poststr. matior	Poststratified ratio esti- mation sample weight by yr			Average August 1967 annual Propor population death rate distrib		
	1966	1967	1968	1966-68	1966	1967	1968	sample deaths, 1966–68*	at-risk (de 100	(deatha/ 100,000), 1 966–68	of 1940 U.S. population
35-44 45-54 55-64 65-74 76-84 Total: 35-84	18 17 17 11 5 68	15 12 10 7 7 51	13 17 18 10 7 65	46 46 45 28 19 184	81 163 313 639 639	82 162 319 640 640	84 166 314 611 611	8,780 7,537 14,163 17,619 11,952 55,051	5,301,563 5,514,745 5,496,482 4,847,416 2,785,516, 23,945,722	- 23.8 45.6 85.9 121.2 143.0 59.4	0.345440 0.292283 0.199205 0.120142 0.042930 1.000000

Weight No. of sample deaths in each 10-yr age group is obtained by multiplication: Unweighted No. of sample deaths in each death year times the poststratified ratio estimation sample weight for the same death year; add results for the 8-yr total. Age-adjusted death rate is calculated by the direct method with the 1940 U.S. population 85-84 yr of age as the standard.

certificate) is determined from information provided by physicians and coded at the NCHS according to the rules of the Seventh Revision of the ICD (7, 8).

Current Population Survey

The CPS (9) is the source of estimates of the non-smoking and smoking populations. Smoking information was collected by the U.S. Bureau of the Census in its CPS of August 1967. Designed to be a representative sample of the Nation's noninstitutionalized population, the total consisted of about 50,000 households (residents of military bases, homes for the sick and aged, and other institutions were excluded). The overall response rate was 95%; of the 95,532 adults in the sample, those between the ages of 35 and 84 years numbered 60,920 and included 25,266 white males and 29,308 white females. Based on this sample, the age-sex distributions of the noninstitutionalized population according to smoking status are shown in Appendix

table 3. Also shown for comparison is the CPS for July I, 1967 (12), which represents the total resident population.

The U.S. Bureau of the Census interviewers recorded information about household and personal characteristics of those in the sample. Demographic and socioeconomic information for all persons in the household was reported by a single adult respondent. Smoking information was obtained from each adult responding for himself, if this was possible during the one interview. However, the smoking information for ! not at home was also obtained from the person incorviewed for the household. Inasmuch as men are more likely than women to be away from home when the interviews take place, only about one-third of the men, as compared with three-fourths of the women in the sample, reported on their smoking habits. Smoking information was missing for only about 3% of the sample.

Estimates of population frequencies of nonsmokers by demographic characteristics were obtained from

APPENDIX TABLE 2.—Estimated 1966-68 U.S. total deaths of whites based on representative sample of deaths from the NMS by age, sex, and cigarette smoking status

Age, yr		ho never smoked rettes*	U.S. whites who cigar	currently smoke	All U.S. whites		
	Male	Female	Male	Female	Male	Female	
35-39 40-44 45-49 50-54 55-59 60-64 65-69 70-74 75-79 80-84 Total: 35-84	6,700 (81) 8,521 (105) 13,734 (93) 25,868 (170) 39,723 (172) 68,025 (265) 90,047 (273) 109,536 (397) 168,226 (498) 157,309 (478) 667,689 (2,532)	7,840 (98) 16,285 (200) 20,697 (132) 84,463 (222) 47,388 (179) 79,733 (309) 131,184 (304) 217,599 (515) 290,082 (708) 298,569 (691) 1,143,840 (8,358)	28,181 (346) 49,359 (608) 74,458 (528) 118,466 (850) 146,149 (774) 168,468 (909) 153,307 (691) 144,226 (656) 113,174 (482) 70,960 (270) 1,066,748 (6,114)	12,067 (148) 21,116 (262) 33,045 (221) 39,771 (273) 48,336 (206) 46,688 (209) 48,520 (139) 41,403 (113) 39,244 (96) 20,659 (53) 350,849 (1,720)	38.858 (476) 66.804 (824) 103.119 (732) 169.694 (1,207) 235.209 (1,193) 293.031 (1,527) 327,846 (1,351) 363,764 (1,547) 377,154 (1,426) 309.449 (1,035) 2,284,928 (11,318)	22,600 (279) 42,220 (521) 60,191 (398) 84,686 (566) 110,514 (448) 141,862 (589) 195,314 (488) 280,660 (698) 349,985 (862) 337,428 (787) 1,625,460 (5,636)	

The first No. in each column is the weighted No. of deaths with the use of weight based on the poststratified ratio estimation procedure. Numbers in parentheses are the unweighted No. of sample deaths used in the NMS.

APPENDIX TABLE 8.—Estimated 1987 age distribution of U.S. whites 25-84 years old by eigarette smoking status based on the CPS

Age, yr		es who never cigarettes		who currently igarettes		All U.S. whites				
	Male*	Female*	Male*	Female*	Male (August 1967)	Male (July 1, 1967)	Female (August 1967) ^e	Female (July 1, 1967)		
85-39	1,168,421	2,527,542	2,779,025	2,205,703	4,779,120	5,002,000	5.137.756	5,172,000		
40-44	1,185,151	2,774,021	8,064,389	2,307,515	5,235,199	5,378,000	5,588,969	5,627,000		
45-49 .	1,195,262	2.788.674	2,894,887	2,248,005	5,054,911	5,161,000	5,434,824	5,464,000		
50-54	1.119.008	2,726,071	2,541,111	1,819,272	4,647,741	4,723,000	4,965,615	5,001,000		
55-59	1,009,872	2,778,288	2,084,528	1.350.491	4.095,290	4.153,000	4,454,802	4,489,000		
60-64	923,158	2.718.194	1,559,549	843,503	3,392,068	8,452,000	8,835,309	8,875,000		
65-69	937,471	2,564,226	956,542	484,001	2,663,209	2,714,000	8,221,374	8,269,000		
70-74	883,178	2.283,190	552,483	241,868	2,009,687	2,067,000	2,650,990	2,726,000		
75-79	695,297	1.834,020	326,279	107,273	1,439,539	1,462,000	2,002,833	2,038,000		
80-84	408,598	951,496	114.336	44,890	689,562	800,000	1,012,735	1.195.000		
Total: 35-84	9,525,416	23,945,722	16,873,129	11,652,521	84,006,326	84,912,000	88,304,207	8 8,856,000		

Numbers are from the August 1967 CPS of noninstitutionalized U.S. population by cigarette smoking status (9).4

Numbers are from the July 1, 1967 CPS of total resident U.S. population (12).

inflation factors supplied by the U.S. Bureau of the Census in a magnetic tape of the sample person records. The inflation factors take into account the sampling fraction and demographic characteristics of those in noninterviewed sample households. Missing demographic characteristics of those in interviewed households were computer assigned except for smoking information. "Not stated" or otherwise unknown smoking status is distributed statistically in proportion to the relative frequencies in the known smoking categories. This procedure is followed, for lack of evidence of a smoking-related pattern of nonresponse, so the best estimate of population can be provided. Because it incorporates U.S. census control totals, the CPS is representative of geographic units, age, sex, and other demographic characteristics of the national population.

Standard Death Rates

Expected numbers of deaths in table 3 are the product of a set of standard death rates for 1966-68 (Appendix table 2) and the August 1967 CPS estimates of whites who never smoked cigarettes (Appendix table 3). The standard rates reflect the mortality experience of the total white population (irrespective of smoking history). These rates are specific for 5-year age groups, sex, and the underlying causes of cancer death shown in table 3. We computed the standard rates by dividing the total of the NMS death estimates for 1966, 1967, and 1968 by the CPS population estimate for August 1967. Thus both expected and observed deaths were derived from the same data base of deaths. In other words, for each cause of death, the total number of expected deaths equals the total number of observed deaths for all white males and females. This procedure minimizes some biases that otherwise may be present in the SMR when the expected deaths are based on an arbitrary standard such as the 1970 U.S. white population.

Sources of Error

Various types of systematic errors likely to affect the estimates of mortality from smoking are discussed below. Also, the reliability of the estimated statistics in relation to sampling error and other sources of random variability is evaluated.

1) Matching between deaths and populations: The major source of bias affecting the results of this study arises from the cross-sectional two-sample design. Theoretically, the decedents are a subset of the total study population; ideally, the prospective design best approximates this concept. The American Cancer Society and U.S. Veterans cohorts (10, 11) are prospective studies in which a beginning sample is initially questioned regarding current and past smoking habits. The cohort is followed longitudinally, and its decedents are recorded with respect to length of time in the study. Death rates are computed as follows: the total number of deaths divided by the total number of person-years lived during the study period by persons in various subgroups. These persons are classified by smoking habits and other characteristics. The decedents of a prospective study, included in the original sample, match exactly with an individual in the study popu-

The present cross-sectional study, however, must rely on statistical matching between the two independently drawn samples: decedents and population. This problem is common to conventional vital statistics rates (7, 8) and must be balanced with the expense, time, and follow-up problems of prospective studies.

Because mortality risk is measured by ratio of deaths to population, the matching problem is a basic consideration in the evaluation of errors from sampling. the data-collection process, and measurement techniques.

2) The samples: A comparability problem exists in the extent to which the target populations differentially represent their universe. The universe of all

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deaths occurring to persons 35-84 years old in the United States during 1966-68 is represented by official registrations of deaths occurring in the United States in any of the 3 years, irrespective of month of death or usual residence of the deceased. However, the population interviewed in the CPS consists of individuals who were not residents of institutions in the United States as of August 1967; this count was at about the midpoint and is believed to represent the 1966-68 population adequately. The CPS practice of replacing part of the sample every 4 months (rather than monthly) extends parts of the target population to cover a span of 7 months, May to November 1967.

Conversely, exclusion of residents of institutions introduces a systematic bias that understates the population and thereby inflates death rates. A first consideration is also to exclude deaths of residents of institutions. The rationale for not doing so is as follows: a) In terms of death rates, the bias is really a function of person-years of excluded institutional population. Fragmentary evidence suggests that many persons in the institutional population are there for only a fraction of a year. b) There is no sound basis for the estimation of the person-years of excluded institutionalized population for the various socioeconomic groups classified by smoking status. c) The net effects of the bias are likely to be canceled in measures of differential mortality like the SMR. d) The bias is most likely to affect only the extremely old, a small fraction of the study population for whom differential mortality is greatly reduced and not of major interest. Only for persons 80-84 years of age is the excluded population more than 5% of the total (Appendix table 3).

3) The respondents: Unlike sizable proportions of the population sample who could provide information on their smoking habits and other personal characteristics, no such data were available from the death sample. Moreover, a mailed questionnaire was used in the death survey, whereas a household interview was conducted for collection of data on the living population. Other sources report substantial agreement between the smoking habits reported for a decedent retrospectively and his reports as obtained in a household interview prior to death (2). With regard to social and demographic characteristics, evidence indicates agreement between the 1960 NMS responses and the matched 1960 census reports (13).

The overall response rates to the NMS and the CPS were nearly equal at 95%; these high levels eliminate any serious bias from differential nonresponse. There is no evidence on nonresponse differentials by smoking status from this investigation. In retrospective studies, in which nonresponse may be as much as 32% overall and even more among smokers, the bias in the mortality ratio of the respondent has been estimated to be about a 0.2 or 0.3 overstatement because of lower mortality ratios of the nonrespondents (1). Again, the low level of nonresponse in the present study (5%) indicates that nonresponse is negligible as a source of bias.

4) Measurement of cigarette smoking: The operational definitions of lifetime history concepts of never smoking cigarettes are almost exactly the same in death survey (<5 packs of cigarettes) and population survey (<100 cigarettes). Copies of the actual NMS and CPS questionnaires are shown elsewhere. Although the slight variations in these definitions do not appear to be a problem, a comparability problem does arise from the fact that those reporting for the decedent tend to understate cigarette consumption compared with what the decedent would have reported as his habits. Haenszel et al. (2) noted that this bias is largest with respect to medium and heavy smokers but virtually nonexistent for nonsmokers.

5) Sampling and data processing: Complex sampling design, coding of reported information, construction of data files, and production of the final data tables are all possible sources of systematic error. The assumption is that these errors are negligible. Table 1 and Appendix table 3 show that the study estimates of death rates and population-at-risk are representative of the United States for the study period. Table 1 illustrates the good agreement between the total age-specific death rates as calculated in the survey and the 1966-68 U.S. vital statistics (8). The good agreement between the total age-specific U.S. white population as determined by the August 1967 and the July 1, 1967, CPS (12), which was used to calculate the 1967 U.S. death rates (8), is depicted in Appendix table 3.

6) Reliability: In addition to systematic error, the estimates reported in this study contain some randor error. This kind of variability is inherent in the sampling of deaths and population at a particular time. In more detailed analyses, random variation may be present in the indicators of mortality that are used for empirical tests of hypotheses. The 1964 Surgeon General's Report evaluated the stability of the SMR by a main assumption that deaths occurring within a given period are generated by a Poisson process (1). The random variability of SMR was small and depended on the number of deaths. The analysis was of 26,000 deaths in 7 prospective studies. Inasmuch as this study is based on about 20,000 sample deaths, one can assume that variability in its SMR is a negligible problem for the major causes of death.

7) Sampling error: In addition to variability over time, the SMR may contain some random variation due to sampling. Both death and population estimates are based on complex, multistage sample designs. Another analysis of the 1967 CPS shows that simple random sampling standard errors understate the true standard errors of estimated population statistics (14). Standard tests of statistical significance are therefore inapplicable to this study.

Hammond (10) used a frequency of 10 deaths as a minimum for standards of reliability. Although arbitrary, the same rule of thumb (adopted in this study) primarily affects deaths when classified by detailed underlying causes. The sample numbers of deaths from the 1966-68 NMS are presented in tables 1-3.

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8) Response error: At the level of the individual, reported information may vary from one observation to another. This variation may introduce error in estimates of fixed characteristics like a history of ever smoking. Kirchoff and Rigdon (15) compared the smoking habits of a group of Texans and reported

different smoking habits for the same reference period when respondents were reinterviewed. This difference partly reflects some memory lapse and some response error in smoking information. We assume that the extent to which this source of variation affects estimates of the present study is negligible but unknown.